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 - due April 7 @ 14:00 in SC1 241

Using the method of maximum likelihood, showed that the best-fit slope & intercept for a line through a linear dataset $(x_i, y_i \pm \sigma_i)$ for $i = 1 \dots N$ are given by:

$$a = \frac{1}{\Delta} \left(\sum \frac{x_i^2}{\sigma_i^2} \sum \frac{y_i}{\sigma_i^2} - \sum \frac{x_i}{\sigma_i^2} \sum \frac{x_i y_i}{\sigma_i^2} \right) \quad \text{intercept}$$

$$\Delta = \sum \frac{1}{\sigma_i^2} \sum \frac{x_i^2}{\sigma_i^2} - \left(\sum \frac{x_i}{\sigma_i^2} \right)^2$$

Likewise:

$$b = \frac{1}{\Delta} \left(\sum \frac{1}{\sigma_i^2} \sum \frac{x_i y_i}{\sigma_i^2} - \sum \frac{x_i}{\sigma_i^2} \sum \frac{y_i}{\sigma_i^2} \right) \quad \text{slope.}$$

Today: Apply prop. of errors to find σ_a & σ_b
 $\Rightarrow$ the uncertainties in the intercept & slope.

Recall: $\frac{\partial y_i}{\partial y_j} = \delta_{ij} = \begin{cases} 1 & i=j \\ 0 & i \neq j \end{cases}$

and $\sum_{i=1}^N f(x_i) \delta_{ij} = f(x_j)$

$$= f(x_1) \cancel{\delta_{1j}} + f(x_2) \cancel{\delta_{2j}} + \dots + \underbrace{f(x_j) \delta_{jj}}_{i=j \text{ term}} + \dots + f(x_N) \cancel{\delta_{Nj}}$$

$= f(x_j)$

What are the uncertainties in a & b ?

$$a = a(y_1, y_2, \dots, y_N)$$

$$b = b(y_1, y_2, \dots, y_N)$$

$$\sigma_a^2 = \sum_{j=1}^N \left(\frac{\partial a}{\partial y_j} \sigma_j \right)^2 \quad (1)$$

$$\sigma_b^2 = \sum_{j=1}^N \left(\frac{\partial b}{\partial y_j} \sigma_j \right)^2$$

First, note that Δ is indep. of y_i

Start by evaluating:

$$\frac{\partial a}{\partial y_j} = \frac{1}{\Delta} \frac{\partial}{\partial y_j} \left(\sum \frac{x_i^2}{\sigma_i^2} \sum \frac{y_i}{\sigma_i^2} - \sum \frac{x_i}{\sigma_i^2} \sum \frac{x_i y_i}{\sigma_i^2} \right)$$

$$= \frac{1}{\Delta} \left(\sum \frac{x_i^2}{\sigma_i^2} \sum \frac{1}{\sigma_i^2} \underbrace{\frac{\partial y_i}{\partial y_j}}_{\delta_{ij}} - \sum \frac{x_i}{\sigma_i^2} \sum \frac{x_i}{\sigma_i^2} \underbrace{\frac{\partial y_i}{\partial y_j}}_{\delta_{ij}} \right)$$

$$= \frac{1}{\Delta} \left(\sum \frac{x_i^2}{\sigma_i^2} \underbrace{\left[\frac{1}{\sigma_i^2} \delta_{ij} \right]}_{\frac{1}{\sigma_j^2}} - \sum \frac{x_i}{\sigma_i^2} \underbrace{\left[\frac{x_i}{\sigma_i^2} \delta_{ij} \right]}_{\frac{x_j}{\sigma_j^2}} \right)$$

$$\frac{\partial a}{\partial y_j} = \frac{1}{\Delta} \left(\frac{1}{\sigma_j^2} \sum \frac{x_i^2}{\sigma_i^2} - \frac{x_j}{\sigma_j^2} \sum \frac{x_i}{\sigma_i^2} \right)$$

Sub $\frac{\partial a}{\partial y_j}$ back into ①:

$$\sigma_a^2 = \sum_{j=1}^N \left(\frac{\partial a}{\partial y_j} \sigma_j \right)^2$$

$$= \frac{1}{\Delta^2} \sum_{j=1}^N \left(\frac{1}{\sigma_j} \sum_i \frac{x_i^2}{\sigma_i^2} - \frac{x_j}{\sigma_j} \sum_i \frac{x_i}{\sigma_i^2} \right)^2$$

Expand the square:

$$\sigma_a^2 = \frac{1}{\Delta^2} \sum_{j=1}^N \left[\frac{1}{\sigma_j^2} \left(\sum_i \frac{x_i^2}{\sigma_i^2} \right)^2 + \frac{x_j^2}{\sigma_j^2} \left(\sum_i \frac{x_i}{\sigma_i^2} \right)^2 - 2 \frac{x_j}{\sigma_j^2} \sum_i \frac{x_i^2}{\sigma_i^2} \sum_i \frac{x_i}{\sigma_i^2} \right]$$

Now, distribute the sum over j .

$$\sigma_a^2 = \frac{1}{\Delta^2} \left[\sum \frac{1}{\sigma_j^2} \left(\sum \frac{x_i^2}{\sigma_i^2} \right)^2 + \sum \frac{x_j^2}{\sigma_j^2} \left(\sum \frac{x_i}{\sigma_i^2} \right)^2 - 2 \underbrace{\sum_{j=1}^N \frac{x_j}{\sigma_j^2}}_{\left(\sum \frac{x_i}{\sigma_i^2} \right)^2} \underbrace{\sum_{i=1}^N \frac{x_i}{\sigma_i^2}}_{\left(\sum \frac{x_i}{\sigma_i^2} \right)^2} \right]$$

$$\sigma_a^2 = \frac{1}{\Delta^2} \left[\sum \frac{1}{\sigma_j^2} \left(\sum \frac{x_i^2}{\sigma_i^2} \right)^2 + \cancel{\sum \frac{x_j^2}{\sigma_j^2} \left(\sum \frac{x_i}{\sigma_i^2} \right)^2} - \cancel{2 \sum \frac{x_j^2}{\sigma_j^2} \left(\sum \frac{x_i}{\sigma_i^2} \right)^2} \right]$$

$$\sigma_a^2 = \frac{1}{\Delta^2} \left[\sum \frac{1}{\sigma_j^2} \left(\sum \frac{x_i^2}{\sigma_i^2} \right)^2 - \sum \frac{x_i^2}{\sigma_i^2} \left(\sum \frac{x_i}{\sigma_i^2} \right)^2 \right]$$

common factor

$$\sigma_a^2 = \frac{1}{\Delta^2} \sum \frac{x_i^2}{\sigma_i^2} \left[\sum \frac{1}{\sigma_j^2} \sum \frac{x_i^2}{\sigma_i^2} - \left(\sum \frac{x_i}{\sigma_i^2} \right)^2 \right]$$

$$\therefore \sigma_a^2 = \frac{1}{\Delta} \sum \frac{x_i^2}{\sigma_i^2}$$

Uncertainty in the best-fit value for the intercept.

Using the same methods, find uncertainty in the slope is given by:

$$\sigma_b^2 = \sum_{j=1}^N \left(\frac{\partial b}{\partial y_j} \sigma_j \right)^2$$



$$\sigma_b^2 = \frac{1}{\Delta} \sum \frac{1}{\sigma_i^2}$$

uncertainty in the best-fit value of the slope.

Last time we had:

$$\underbrace{\sum \frac{y_i}{\sigma_i^2}}_u = a \underbrace{\sum \frac{1}{\sigma_i^2}}_{u_a} + b \underbrace{\sum \frac{x_i}{\sigma_i^2}}_{u_b}$$

$$\underbrace{\sum \frac{x_i y_i}{\sigma_i^2}}_v = a \underbrace{\sum \frac{x_i}{\sigma_i^2}}_{v_a} + b \underbrace{\sum \frac{x_i^2}{\sigma_i^2}}_{v_b}$$

Solve system of two eq'ns for 2 unknowns a & b .
Of the form:

$$\begin{pmatrix} u \\ v \end{pmatrix} = \underbrace{\begin{pmatrix} u_a & u_b \\ v_a & v_b \end{pmatrix}}_{\equiv A} \begin{pmatrix} a \\ b \end{pmatrix}$$

$$\therefore \begin{pmatrix} a \\ b \end{pmatrix} = A^{-1} \begin{pmatrix} u \\ v \end{pmatrix}$$

$$\begin{pmatrix} a \\ b \end{pmatrix} = \frac{1}{\det A} \begin{pmatrix} v_b & -u_b \\ -v_a & u_a \end{pmatrix} \begin{pmatrix} a \\ b \end{pmatrix}$$

A^{-1}

Recall, we defined $\det A$ to be Δ .

$$\therefore \begin{pmatrix} a \\ b \end{pmatrix} = \frac{1}{\Delta} \begin{pmatrix} v_b & -u_b \\ -v_a & u_a \end{pmatrix} \begin{pmatrix} u \\ v \end{pmatrix}$$

called the
covariance matrix.

For linear fits, the covariance matrix is given by:

$$\frac{1}{\Delta} \begin{pmatrix} \sum \frac{x_i^2}{\sigma_i^2} & -\sum \frac{x_i}{\sigma_i^2} \\ -\sum \frac{x_i}{\sigma_i^2} & \sum \frac{1}{\sigma_i^2} \end{pmatrix}$$

The diagonal elements of the covariance matrix give the squares of the uncertainties in the best-fit parameters.

The 11 element gives $\sigma_a^2 = \frac{1}{\Delta} \sum \frac{x_i^2}{\sigma_i^2}$

The 22 element gives $\sigma_b^2 = \frac{1}{\Delta} \sum \frac{1}{\sigma_i^2}$

Linearizing Experimental Data:

If we can analyze data using linear fits, it's good to do b/c we can exactly determine

$a \pm \sigma_a$ (intercept)

$b \pm \sigma_b$ (slope)

This process is easy when a set of meas. is obviously linear.

Eg. Object in free fall



$$V = V_0 - gt$$

of the form

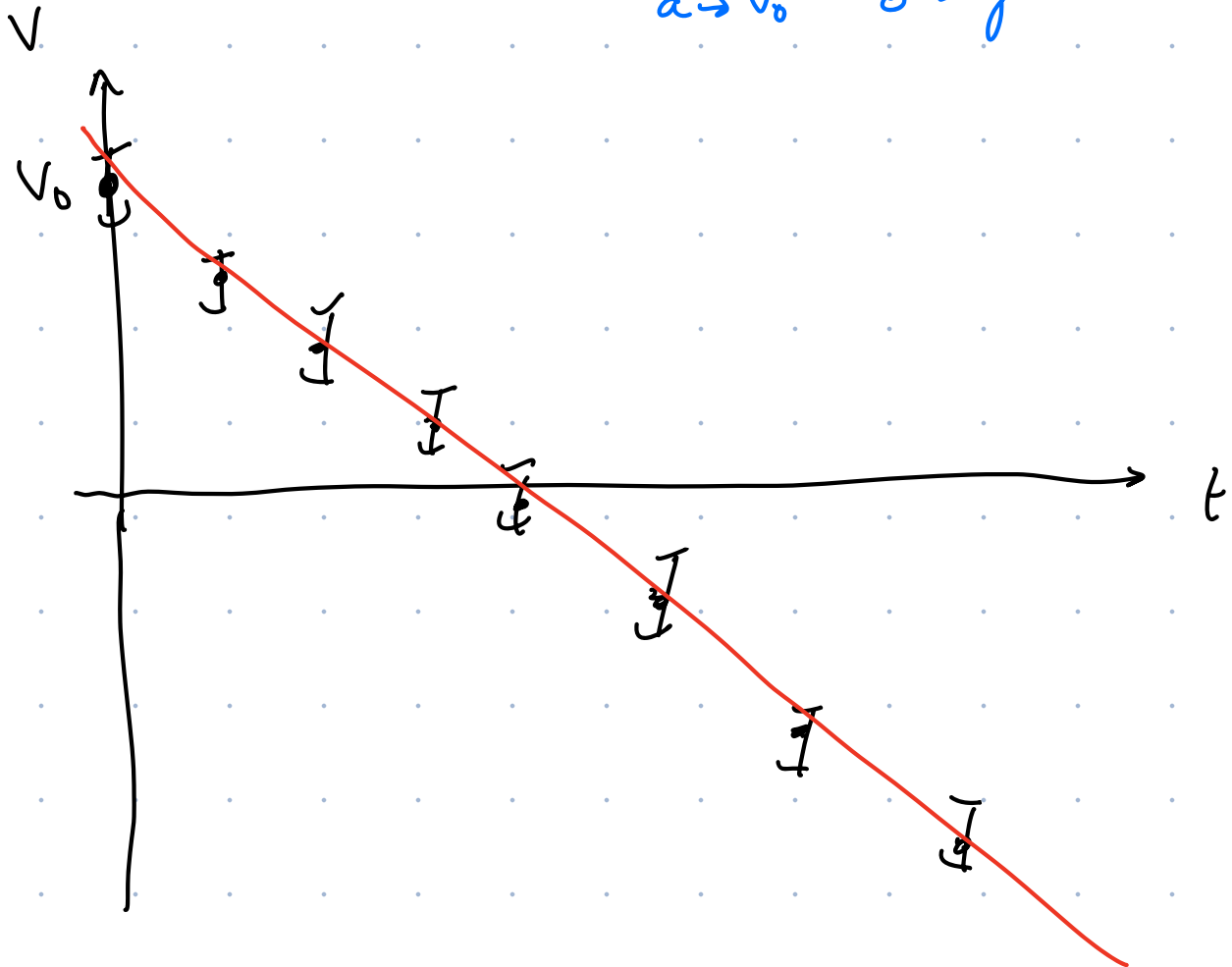
$$y = a + bx$$

$$y \rightarrow V$$

$$x \rightarrow t$$

$$a \rightarrow V_0$$

$$b \rightarrow -g$$



Sometimes, a pair of variable may not be strictly linearly related, but often can still use linear fits to analyze the data.